

Probability & Statistics,

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The probability density of the **gamma distribution** is given by

$$
f(x) = \begin{cases} \frac{1}{\Gamma(\alpha)\beta^{\alpha}} x^{\alpha-1} e^{-x/\beta} & \text{for } x > 0, \ \alpha > 0, \ \beta > 0\\ 0 & \text{elsewhere} \end{cases}
$$

where $\Gamma(\alpha)$ is a value of the **gamma function**, defined by

$$
\Gamma(\alpha) = \int_0^\infty x^{\alpha-1} e^{-x} dx
$$

Properties:

$$
\sum \Gamma(\alpha) = (\alpha - 1)\Gamma(\alpha - 1) \text{ for any } \alpha > 1.
$$

Proof:

$$
\Gamma(\alpha) = -e^{-x} x^{\alpha - 1} \bigg|_0^{\infty} - \int_0^{\infty} [-e^{-x} (\alpha - 1) x^{\alpha - 2}]
$$

$$
= 0 + (\alpha - 1) \int_0^{\infty} e^{-x} x^{\alpha - 2} = (\alpha - 1) \Gamma(\alpha - 1)
$$

$$
\sum \Gamma(1) = 1 \text{ and } \Gamma(1 / 2) = \sqrt{\pi}
$$

$$
\sum \Gamma(\alpha) = (\alpha - 1)! \text{ if } \alpha \text{ is a positive integer}
$$

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Figure: Graph of some gamma probability density functions

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Mean of gamma distribution:

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$$
\mu = \alpha \beta
$$

$$
\frac{\text{Proof:}}{\mu} = \frac{1}{\beta^{\alpha} \Gamma(\alpha)} \int_0^{\infty} x \cdot x^{\alpha-1} e^{-x/\beta} dx
$$
\n
$$
= \frac{\beta}{\Gamma(\alpha)} \int_0^{\infty} y^{\alpha} e^{-y} dy = \frac{\beta \Gamma(\alpha+1)}{\Gamma(\alpha)}
$$
\n
$$
= \alpha \beta
$$
\n(put $y = x/\beta$)

Problem 1

If X_1, X_2, \dots, X_n are independent random with parameters $(\alpha_1, \beta), (\alpha_2, \beta), \cdots, (\alpha_n, \beta)$ respectively. Show that $X_1 + X_2 + \cdots + X_n$ $(\alpha_1 + \alpha_2 + \cdots + \alpha_n, \beta).$ variables and follows gamma distribution follows gamma distribution with parameter

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The density function of exponential distribution is given by

$$
f(x) = \begin{cases} \frac{1}{\beta} e^{-x/\beta} & \text{for } x > 0, \ \beta > 0 \\ 0 & \text{elsewhere} \end{cases}
$$

Note: Exponential is the special case of gamma distribution where $\alpha = 1$.

Mean, variance and generating function of the exponential distribution:

$$
\mu = \beta, \sigma^2 = \beta^2 \text{ and } M_X(t) = \frac{1}{1 - \beta t}
$$

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Problem 2

If X is exponential random variable with parameter β . Find the cumulative distribution function of X.

$$
Ans\colon F(x) = 1 - e^{-x/\beta}
$$

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Problem 3

If X_1, X_2, \dots, X_n are independent exponential distributions with parameters $\beta_1, \beta_2, \dots, \beta_n$ the random variable $Y = min(X_1, X_2, \dots, X_n)$. respectively. Find the density function of

Memory Loss Property

Problem 4

If X is exponential random variable with parameter β . Show that $P(X > x + t | X > t) = P(X > x)$ for $t > 0$.

Note: The converse also true. Verify ?

Memory Loss Property

Remark:

- **►** If *X* represents the lifetime of a device, then memory loss property states that if the device has been working for time *t*, then the probability that it will survive an additional time *x* depends only on *x* (not on *t*) and is identical to the probability of survival for time *x* of a new device.
- \triangleright The equipment does not remember that it has been in use for time *t*.

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Memory Loss Property

Problem 5

Suppose the life length of a machine has an exponential distribution with $\beta = 10$ years. A 7 years used machine is bought by someone. Find the probability that it will not fail in the next 5 years.

Problem 06

Personnel of a company use an online terminal to make routine calculations. If the time each person spends in a session at a terminal has an exponential distribution with an **average value of 36 minutes,** find the probability that a person

- (i) Will spend 30 minutes or less at the terminal
- (ii) If he has already been at the terminal for 30 minutes, what is the probability that he will spend more than another hour at the terminal.

Theorem:

If in a Poisson process average number of arrivals per unit time is λ . Let W denote waiting time between successive arrivals (or the time until the first arrival). Then W has an exponential distributi on with 1 $\beta = \frac{1}{a}$. λ $=$

Proof:

Let X be the number of arrivals in the
time interval of length t.

$$
\Rightarrow
$$
 X follows Poisson process and

$$
P(X = x) = \frac{(\lambda t)^x e^{-\lambda t}}{x!}, x = 0, 1, \cdots
$$

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Proof(cont.):

 $\left(\lambda t\right)^{0}e^{-\lambda t}$ P (waiting time between successive arrivals be at least t) $= P(W > t)$ P (no arrivals during a time interval of length t) \Rightarrow *P*(waiting time between successive arrivals $\langle t \rangle$ $= 1 - e^{-\lambda t} = F(t)$, the distribution function of W. 0! $f)^0 e^{-\lambda t}$ *e* $\lambda t)^0 e^{-\lambda t}$ $_{-\lambda}$ $=\frac{(\pi e)^e}{e^+}$ = e⁻

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Proof(cont.):

So if waiting time between successive arrivals be random variable with the distribution function

$$
F(t)=1-e^{-\lambda t}
$$

So, the probability density of the waiting time between successive arrivals given by

$$
f(t) = \frac{d}{dt} F(t) = \lambda e^{-\lambda t},
$$

 which is an exponential distribution with

$$
\beta = \frac{1}{\lambda}.
$$

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Problem 07

Given that the switchboard of consultant's office receives on the average 0.6 calls per minute, find the probabilities that the time between successive calls arriving at the switchboard of the consulting firm will be

(a) less than 1/2 minute;

(b) more than 3 minute.

Solution to Problem 07

Since $\lambda = 0.6$, the waiting time *t* between successive calls arriving at the switchboard, has an exponential distribution with $\beta = 1/0.6$, hence density function is given by

$$
f(t) = \begin{cases} 0.6e^{-0.6t} & \text{for } t > 0\\ 0 & \text{elsewhere} \end{cases}
$$

(a)
$$
P(t < \frac{1}{2}) = \int_{0}^{1/2} 0.6e^{-0.6t} dt = -e^{-0.6t}\Big|_{0}^{1/2} = 1 - e^{-0.3}.
$$

(b)
$$
P(t > 3) = \int_{3}^{\infty} 0.6e^{-0.6t} dt = -e^{-0.6t} \Big|_{3}^{\infty} = e^{-1.8}.
$$

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Problem 08

Customers arrive in a certain shop according to an approximate Poisson process at a mean rate of 20 per hour. What is the probability that the shopkeeper will have to wait more than 5 minutes for the arrival of the first customer? .

Problem 09

An average of 30 customers per hour arrive at a shop in accordance with a Poisson process. Find the probability that the shopkeeper will wait more than 5 minutes before the second customer arrives.

The density function of Chi-squared distribution is given by

$$
f(x) = \begin{cases} \frac{1}{2^{\gamma/2} \Gamma(\gamma/2)} x^{\gamma/2-1} e^{-x/2}, & x > 0 \\ 0, & \text{elsewhere} \end{cases}
$$

We denote the random variable by X_{γ}^2

and γ is called degree of freedom.

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Mean, variance and moment generating function of the Chi-squared distribution:

$$
E(X_{\gamma}^{2}) = \gamma \left[Var(X_{\gamma}^{2}) = 2\gamma\right]
$$

$$
M_{X_{\gamma}^{2}}(t)=(1-2t)^{-\gamma/2},\ t<1/2
$$

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Advantage of Chi-squared distribution

Remark:

- \triangleright Cumulative distribution values available for some points with specified degree of freedom.
- \triangleright Sum of independent Chi-squared distributions is again Chi-squared.

How to use cumulative distribution Table:

Applications

- \triangleright The aggregate insurance claims and the amount of rainfall accumulated in a reservoir are modeled by a gamma distribution.
- \triangleright In neuroscience, the gamma distribution is often used to describe the distribution of inter-spike intervals.
- \triangleright In genomics, the gamma distribution was applied in peak calling step (i.e. in recognition of signal) in ChIP-chip and ChIP-seq data analysis

Applications

- \triangleright In bacterial gene expression, the copy number of a constitutively expresses protein often follows the gamma distribution, where the scale and shape parameter are, respectively, the mean number of bursts per cell cycle and the mean number of protein molecules produced by a single mRNA during its lifetime.
- \triangleright It also used to fit life model of many devices based on certain assumptions.

